Response Letter

Referee 1

-- Ref1.1 - General comments -

Reviewer' comment:

The authors have addressed all my concerns.

Author's response:

We appreciate referee 1's comments.

Referee 2

-- Ref2.1 - General comments -

Reviewer' comment:

The authors are right in stating that there is no other reference that implements a noncoding mutation burden analysis. The only one I know is Weinhold et al., (2014), and I also agree with the authors that a simple binomial test as applied in that reference is not good enough to correctly compute the mutation burden in noncoding regions. What I wonder is if the change from a binomial to a beta-binomial distribution is a good enough solution. Unfortunately, the controls provided in the new version don't seem to be enough to prove that, see comments below. Also, we have tried to run the software and we found many problems and unsatisfactory results in the only case we managed to run it (described below).

Overall I agree with the authors that it would be an important contribution to describe and provide a method that does the noncoding mutation burden analysis correctly. I am not convinced that LARVA does it well, at least in its current version, based on our test on running the software (see below) and on the description in the manuscript.

Author's response:

We thank the reviewer for agreeing with contributions. We have further investigated the false positive and negative problems through simulation and permutation studies. The reviewer's concerns were answered point by point here.

<u>1. "What I wonder is if the change from a binomial to a beta-binomial distribution</u> is a good enough solution"

We added more discussions about importance of <u>covariate</u> correction (paragraph 5 in discussion and section 5 in Text S1) and how to interpret the usage of <u>the</u> beta-binomial model (paragraph 4 in discussion). We actually also tried other distributions, like the negative binomial, <u>and</u> Poisson inverse Gaussian. The <u>performance of these distributions</u> is similar to that of the beta-binomial distribution. Eventually we selected beta-binomial because of <u>its</u> immediate interpretability.

2. "we have tried to run the software and we found many problems and unsatisfactory results in the only case we managed to run it"

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We have addressed the software issues raised by the reviewer. Detailed answers <u>are</u> provided in section Ref2.5.

-- Ref2.2 - False positive and false negative rate -

Reviewer' comment:

AUTHOR'S RESPONSE We emphasize our contribution in the following listed points. 1. We are among the first to implement the somatic burden test with overdispersion control, which is specifically designed for noncoding somatic variant analysis. MY NEW COMMENT I agree with that. It is important not only to be among the first but more importantly to make sure that the test is correct, give a good control of false positives and false negatives, and provide a code that users can run. AUTHOR'S RESPONSE 2. We release a convenient annotation resource for the whole community by gathering all the noncoding regulatory regions from more than 122 experiments from the ENCODE project. Notably, this data has never been collected in one place before, which will greatly facilitate subsequent research. 3. Our released noncoding regulatory element corpus provides a natural and meaningful solution about how to pool biologically relevant regions to perform the mutation burden test. We do not have to rely on the bin procedure, which is a relatively ad-hoc method. 4. Once highly mutated regions are detected in a certain cancer type, users can immediately understand the functions of these regions. MY NEW COMMENT I agree with authors that 2, 3 and 4 are useful additional resources provided with the code of LARVA, however the first and more important

provided with the code of LARVA, however the first and more important think is that authors convince that LARVA is able to detect noncoding recurrently mutated drivers, which I understand from the description of the paper it is the main aim, with an acceptable rate of false positives and false negatives. This is not clear in this version of the software and manuscript.

<u>Author's response:</u>

We thank the reviewer for agreeing with our contribution. We added some simulation and permutation studies for these questions.

1. Simulation studies

- For all 2.5kb bins on the genome, remove those intersecting gap regions and blacklist regions. For the remaining 1,139,452 bins, count the variants.
- Use negative binomial regression to build <u>the</u> mutation model NB(μ_i, σ_i) for the ith bin by correcting replication timing, GC content, and chromatin status.
- Simulate the variant counts in 1,139,452 bins using local NB(μ_i, σ_i)



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 Randomly selected 100 bins as the true signal, replace the variant count as the top 1e-4 quantile of local NB(10 * μ_i, σ_i)

Figure R1 gives the 3d plot on how the average mutation rate changes with these three covariates.



Then we used beta-binomial and binomial model by only correcting the replication timing effect. Our method gives 108 positives—including all true positives,—and the binomial test gives $48_{a}505$ positives (BH adjusted P<0.05).

2. Permutation studies

For each variant in a set of whole genome sequencing data, find a new position in <u>a</u> 25kb neighboring region (12.5k and 12.5k up and downstream each). Then we tested all the noncoding regions on the original and permuted data set. Since the permuted size 25kb is relatively large as compared to the test region, a better method is supposed to give less or even no positives on the permuted data set. The P values were given in the updated Fig. S13 in Text S1. From the Q-Q plots of the P values, it can be seen that LARVA <u>yields fewer</u> positives in larger regions, such as DHS, TF, Promoter and UTRs. In very small regions such as ultra sensitive and ultra conserved regions, the two methods gives similar performance.

--Ref2.3 - P-values for all genes -

Reviewer' comment:

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Finding only 7 significantly mutated coding genes analyzing 5032 tumors
is a surprising low number. I agree that 6759 significantly mutated
genes with the binomial test is a not an acceptable number of genes,
surely full of false positives. It would be useful if authors provide a
supplementary table with the obtained pvalue per gene, not only for the
7 genes claimed as highly mutated by LARVA.
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Author's response:

We thank the reviewer for this comment.

1). As requested, we provide the P values on our website

2). We agree that 7 is a rather low number but this is understandable since our method is not optimized in the coding region analysis. We clearly mentioned this point in

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Author's response:

We thank the reviewer for bringing this to our attention. We have addressed the long running time by profiling our code, and optimizing the computations in portions of the code where the running time did not scale well with the size of the input. We have released revised code along with our revised manuscript.

Furthermore, we have migrated our R codebase into C^{++} , giving us more direct control over the source code. Our new code is not prone to the error the reviewer encountered.

-- Ref2.6 - Software Errors II -

Reviewer' comment:

We next tried to run LARVA with a dataset of 505 tumor whole-genomes

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across 14 cancer types as reported in Fredriksson et al., 2014 in
promoters and ultra-sensitive regions. For both promoters and ultra-
sensitive regions we used the annotations present in the folder
data/annotations/ of LARVA. The program didn't run successfully on
promoters and raised an error after approx. 12 hours. Following is the
trace of the error:
Error in d$p.bbd.cor[d$p.bbd.cor <= 0] = rep(d$p.tiny, sum(d$p.bbd.cor
<= :
replacement has length zero
Execution halted
```

Author's response:

We thank the reviewer for bring this to our attention. We have determined that this error can occur in rare boundary conditions in our R code. We have migrated our R codebase into C++, and now have more direct control over the functioning of our code. Our new code handles these conditions properly.

-- Ref2.7 - Software P-value Output -

Reviewer' comment:

We finally managed to run LARVA with this dataset in ultra-sensiti [.]	ve	
regions. In this case the program performed the analysis quickly	y.	
However, when we check the files with the results we found cases, wi	th	
the exception of 'p.bbd.cor.adj', where the pvalues were greater than \Im	1.	
How can this be possible? Following are the maximum values of eac	ch	
pvalue type:		
p.bbd 3.488000		
p.binomial 16.4250	00	
9.bbd.cor 3.2860	00	
p.binomial.cor 20.5960	00	
o.bbd.adj 1.1480	00	
o.bbd.cor.adj 0.3570	00	
o.binomial.adj 14.0310	00	
p.binomial.cor.adj 17.464000		

Author's response:

We thank the reviewer for bringing this to our attention. We didn't clearly mention in our software documentation that th<u>ese numbers are</u>, in fact, -log10-transformed p-values, hence the observed output is correct. This has been rectified in the current version's documentation.

-- Ref2.8 - Software P-value QQ Plots -

Reviewer	After filtering the results for regions that overlapped \checkmark	$ \rangle$
Comment	genes or pseudogenes and for regions without mutations, we	$\langle \rangle$
	did QQplots as follow: we discarded pvalues > 1	
	(considering them wrong) and we plot on the y axis the -	
	log10 of the sorted observed pvalues and on the x axis the	
	-log10 of a uniform distribution of expected pvalues	
	between 0 and 1. The QQplots were generated by using the	
	code provided	

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here:http://www.broadinstitute.org/files/shared/diabetes/s candinavs/qqplot.R) The resulting plots showed that the both the 'pbb' pvalues distributions (p.ddb and p.ddb.cor, top row of the figure) are deflated respect to a perfect correlation between observed and expected pvalues (red diagonal line) and thus the methods are finding less significant genes that what expected by the null model. On the other hand the binomial method (bottom row of the figure) is somehow inflated respect to the red diagonal. While the binomial method is likely to find a number of false positive candidates, the method proposed by the authors is likely to miss many true positive candidates.

Author's response:

We thank the reviewer for the careful checking of Q-Q plots. As mentioned in section Ref2.7, our provided P values are already log transformed P values. Hence taking the -log10(P) again does not reflect the real P value here. We plotted the Q-Q plots as suggested by the reviewer in Fig. S13. Our P values follow the uniform distribution line.

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